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**Fatness and overweight in Amerindian women and children from riverine
Amerindian communities of the Beni River (Bolivian Amazon)**

Running head title: Overweight in Amerindian children and women

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Nature of the contributions: EB was responsible for all stages of the study and wrote the first draft of the paper; EB and RL designed the survey; EB and SLM collected dietary, clinical and anthropometric data; SLM and SR performed parasitological analysis; RL was in charge of laboratory analysis and follow-up of subjects.

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Abstract

Obesity is a worldwide public health problem impacting not only industrialized nations but also developing countries. The main objective of this paper was to analyze risk factors for overweight among Amerindian children and their mothers. Data were collected in 15 Amerindian riverside communities from the Beni River (Bolivia). The total sample was 195 mothers and 452 children, 0-15 years of age. Information about family activity and dietary patterns was collected and a clinical examination was performed. Stool samples were collected in children for parasitological screening. Anthropometric measurements, including weight, height, arm circumference and 4 skinfolds were taken. A bio-electrical impedance analysis was performed in mothers. A total of 12.2% (IC_{95%}: 9.1~15.9%) of the children were considered overweight; less than 1% were overtly obese (IOTF criteria). Among their mothers, 35.3% (IC_{95%}: 28.7~42.5%) were overweight (BMI (weight / height²) > 25 kg/m²) and 5% (IC_{95%}: 2.5~5.1%) were obese (BMI > 30 kg/m²). The BMI was moderately related to anthropometric indices of body composition in children, but was highly correlated with fatness in mothers. The risk of overweight was not associated with environmental factors in children. In mothers, there was a significant relationship between BMI, health status and dietary diversity score. There was a moderate association in boys between fatness and the BMI of their mothers ($R^2 = 0.12$, $p < 0.001$), but not in girls. These findings suggest a trend toward accumulation of fat related to possible changes at the economic and agricultural levels, even in remote rural areas.

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Key words: Body mass index, Electric impedance, Skinfold thickness, Dietary diversity, South American Indians, Bolivia

2 Introduction

4 Up until the last decade, research on malnutrition in developing countries focused almost
exclusively on undernutrition. However, about 7% of the world population is now considered to
6 be obese (Seidell 1999). The explosion of so-called “obesity epidemics” and increased awareness
that non-industrialized countries are also affected have prompted numerous investigations. The
8 rise in obesity is worrisome in developing countries since, in addition to its individual clinical
harm, it is likely to represent an unaffordable cost in funding of public health institutions in
10 terms of diagnosis, handling, treatment and rehabilitation of patients (Pena and Bacallao 2002).

12 The situation in developing countries is obscured by the coexistence of overweight or obesity
and undernutrition, especially chronic malnutrition. A recent report analyzed this so-called
14 “double burden” using data from 36 developing countries. Overweight prevalence exceeded
underweight in more than half of the selected countries. Urbanized and economically emergent
16 countries presented the highest overweight/underweight ratio (Mendez and others 2005). In 8
Latin America countries the median prevalence of overweight was 32.4% in urban women. This
18 trend toward increased obesity and overweight rates in Latin America and Caribbean regions was
first documented by Martorell and coworkers using DHS (Demographic and Health Surveys)
20 data (Martorell and others 1998). Obesity (body mass index or BMI $> 30 \text{ kg/m}^2$) ranged from 8-
10% in Latin American women. Overall, 9 to 14.7% of children were overweight (weight-for-
22 height $> 1 \text{ SD}$ above the reference value). This unexpected overweight risk in children was
thereafter confirmed by an analysis of the WHO “Global Database on Child Growth and
24 Malnutrition” (de Onis and Blossner 2000). Using a weight-for-height cut-off of $>2 \text{ SD}$ as the
entry point into obesity, the authors found the highest prevalence rates in Latin America and the
26 Caribbean (4.4%). The overall prevalence was estimated at 3.3%, thereby exceeding the 2.3%
expected prevalence (de Onis and Blossner 2000).

28
A number of recent Latin American surveys have confirmed the rapid emergence of obesity. In
30 Chile, among poverty-stricken populations in the capital city, there was an increase in
prevalence from 6 to 11% in men and from 14 to 24% in women over only a 4-year period.
32 Overweight also increased in children, especially among the beneficiaries of a food supplement
program (Uauy and others 2001). In Mexico, children have been strongly affected by the obesity

wave: 10 to 16% of boys and 14 to 19% of girls are considered to be overweight (del Rio-
2 Navarro and others, 2004). In Brazil, based on 3 successive national surveys performed in 1975,
1989 and 1996, a rapid increase in obesity was clearly demonstrated. All population groups were
4 affected (Monteiro and others 2000). In large countries such as Mexico and Brazil, considerable
heterogeneity was found according to the geographic location and the economic situation
6 (Abrantes and others 2002; del Rio-Navarro and others 2004).

8 Trends in obesity patterns in Latin America cannot be viewed as a simple linear projection of the
trend in industrialized countries. Substantial heterogeneity exists in Latin American countries
10 from the standpoint of economic development, education and health services, the origins of the
populations (European versus Amerindian) etc., that may involve differences in nutritional
12 status. Bolivia, one of the most poorly developed Latin American countries, had the highest
prevalence of malnutrition among Andean countries (Larrea and Freire 2002). However, recent
14 studies indicate that Bolivia is also experiencing an increase in the prevalence of overweight and
obesity (Amigo 2003; Martorell and others 1998). Recent analyses of DHS surveys point to a
16 rise in an overweight prevalence (defined as $25 < \text{BMI} < 30 \text{ kg/m}^2$) of 9 percent between 1994 and
1998, but not in obesity ($\text{BMI} > 30 \text{ kg/m}^2$). One explanation is that Bolivia is just at the beginning
18 of the “nutritional transition” process (Perez-Cueto and Kolsteren 2004; Popkin 2001). The latest
DHS survey (2003) confirmed this increase in the prevalence of both overweight (30.3%) and
20 obesity (15.1%) in women (MEASURE 2004). This trend is even more pronounced in the capital
city La Paz, and involves school-age children and adolescents (Perez-Cueto and others 2005).
22 The prevalence of overweight in women is higher in urban and lowland areas (“Llano” and
Amazon) than in rural or Andean areas of Bolivia (MEASURE 2004).

24 Since obesity and overweight indices are calculated from height and weight, one important
26 concern is their significance in populations of short stature and with a high degree of chronic
malnutrition, such as those found in Bolivia. According to the latest DHS survey, the mean
28 stature of women is 152 cm in Bolivia; 10.3% of females measured less than 145 cm
(MEASURE 2004). It has been argued that, for the same BMI, subjects with short stature are
30 fatter than taller persons (Lopez-Alvarenga and others 2003). This would mean that the
overweight prevalence in Bolivia is even worse. Likewise, an association between short stature
32 and obesity was found in Brazil, and was more pronounced in women than in men (Sichieri and
others 2003). It is thus important to use independent indicators of fatness, like skinfold thickness
34 and bioelectric impedance analysis. Interestingly, measurement of several skinfolds enabled the

elaboration of a simple index of regional fat distribution, such as the trunk-to-extremities ratio
(Malina and others 1999; Rolland-Cachera and others 1990). Moreover, fat distribution,
especially higher fat accumulation on the trunk in early adolescence, is an acknowledged risk
factor for metabolic diseases during adulthood (Malina and others 1999).

Hence, the main objective of this paper was to present data on overweight and obesity in children
and women in a population leading a “traditional” way of life in the Beni department (lowland
Bolivia). In addition, we sought to elucidate the role of environmental, familial and individual
factors in the risk of overweight.

Methods

Milieu and populations

Results presented here are drawn from a health study performed in 2004 and 2005 in a population living along the Beni River (Bolivia), from the foothills (Rurrenabaque: south =14.4°; west= 67.5°) of the Andes to 120 km downstream (Villa Fatima: south =13.8°; west= 67.5°). The main purpose of that survey was to provide first-hand data on the health and nutritional status of 15 communities exposed to mercury contamination (see map, Figure 1). The study was carried out below the 200 m contour in an area with a humid tropical climate characterized by high annual rainfall, over 2500 mm, and an average temperature of 22 to 24 °C.

The total population living in the area, based on the 2001 census (INE 2001), is estimated at about 14,000, but only 4,000 persons live along the banks of the Beni River. The sample size was chosen on the basis of a reported prevalence of 35 % chronic malnutrition in 0-10-year-old children in rural Bolivia (MEASURE 2004). A theoretical calculation showed that at least 364 children were required in order to obtain 5% precision, with a 95% confidence interval (CI_{95%}). A total of 348 children 0 to 10-years-old were finally measured, providing adequate precision. Communities were constituted by small hamlets scattered along the river, with frequent displacement for fishing, logging and agricultural activities. It was not possible to draw a random sample. Moreover, this procedure would have been misunderstood and unacceptable to the population. We therefore decided to incorporate all mothers and their children up to 15 years of age present in the communities at the time of our visits. Data were collected in 2004 during 4 field rounds (March to August) and completed by a survey carried out in March 2005 in order to verify and complete certain data. The present sample (n=647) comprised 452 children (171 under 5 years of age, 177 between 5 and 10 years of age, and 104 adolescents older than 10 years), along with 195 mothers (33.5 ± 13.1 years old). Anthropometric data (skinfold measurement) were incomplete for 2 children and 4 mothers; hence numbers varied in the table.

Most of the inhabitants belonged to the Takana ethnic group. The others were Esse Ejjas, a closely related group. A small number of families were Aymaras who migrated from the Altiplano to settle in the Beni. Takanas constitute a well-integrated group in modern Bolivian

society and speak Spanish. Wood-cutting is their main wage-earning activity. Esse Ejjas lead a more traditional way of life, subsisting mostly on fishing. While this report does not intend to compare the two groups, there exist differences in their way of life that will be accounted for in this analysis. The living conditions of all riverside communities are unfavorable. Public infrastructures such as health centers, elementary and secondary schools are rare. There exists only limited access by roads, and traveling by canoes or small boats is the means of transportation adopted by most communities. Nutritional and health indicators are mediocre. Subjects have been exposed to multiple infections and parasitic infestations. Chronic malnutrition, represented by stunting in growth, affects 41% of preschoolers; however, no cases of severe malnutrition (wasting) were found. Mothers were found to be of short stature. Iron deficiency anemia is quite frequent: 42% of the women had a hemoglobin concentration under 12 g/100 ml. More than half of the mothers surveyed had lost at least one child (Benefice and others 2006). Food consumption was also analyzed in detail. The main meal was an association of rice and meat or fish. Yucca or *Manihot esculenta*) and plátano (*Musa paradisiaca*) were consumed almost every day by 2/3 of the households. Fruits consumption varied according to the season with a predominance of oranges and grapefruits (Benefice and others 2006).

Ethical considerations

The study protocol was approved by the Facultad de Bioquímica y Farmacia (Universidad Mayor de San Andrés, La Paz), the Seladis Institute (Servicio de Laboratorio de Diagnóstico e Investigación en Salud) and the Institut de Recherche pour le Développement (IRD, France) Review Boards. The Comité Nacional de Bioética (Bolivia) as well as the Comité Consultatif de Déontologie et Bioéthique (CCDE, France) gave their agreement. Subjects were individually notified about the aims of the study and had to sign an informed consent form. In the case of children, their parents signed the form. When subjects were illiterate, a reliable witness signed instead. Subjects freely agreed to participate in the study and did not receive any salary or compensation for this. However, when they were found to be ill, they received free medical orientation and treatment from one of the physicians on the team. Children were offered anti-helminth treatment when they had contaminated stools, and mothers received a 3-month supply of ferrous iron and folate when their hemoglobin concentrations were low.

2 *Study variables*

4 Each subject underwent clinical examination by a physician (EB). A total of 373 subjects
(57.6%) were found to be free of clinically detectable disease; 250 (38.6%) presented mild
6 symptoms and 24 (3.7%) more severe illness requiring medical treatment. For simplicity of
analysis, the health status was grouped into 2 categories (code: 1 = absence of clinical symptoms;
8 2= presence of clinical symptoms). The most frequently observed clinical abnormality was
dermatosis: 30% of the children suffered from pyodermatitis, scabies, impetigo or lice. Common
10 infections such as conjunctivitis, pharyngitis and rhinitis were also frequent (about 20%). Simple
diarrhea was frequently alleged by mothers (20% of responses). Finally more severe conditions
12 (otitis media, pneumonia) were observed in only a few occasions.

14 An anthropometric examination was next carried out (SL & RL). It consisted of measurements
of:

- 16 1) Body weight (kg) recorded with an electronic Seca ® scale (accuracy, 100 g); children
less than 2-years-old were weighed on an electronic Seca ® scale accurate to 10 g.
18 Subjects were weighed lightly clothed.
- 20 2) Standing height (cm) was measured with a Harpenden ® anthropometer (accuracy 1
mm). Children younger than 2 years were measured in a recumbent position with a
wooden board.
- 22 3) Mid-arm circumference (cm) was measured on the left side with an unstretchable tape.
- 24 4) 4 subcutaneous skinfolds (mm) were taken at the triceps (Tric), biceps (Bic), subscapular
(Sca) and supra-iliac (Sil) sites using a Holtain ® caliper. Sites were chosen in
accordance with the recommended technique and marked on the skin before the
26 measurement (Lohman and others 1988).

28 All measurements with the exception of body weight were taken on the right side, in duplicate,
and the average was used in the analysis.

30

Several anthropometric indices were calculated:

32

- 1) Indices of malnutrition according to the World Health Organization (WHO 1995): height

for age (H-age); weight for age (W-age); and height for height (H-W). Anthro software enabled calculation of z-scores (Centers for Diseases Control: <http://www.cdc.gov>).

- 2) The arm muscle circumference (AMC, cm), considered as an indicator of muscularity, was calculated using the formula by Gurney and Jelliffe (Gurney and Jelliffe 1973):

$AMC = \text{arm circumference} - \pi (\text{triceps skinfold}).$

- 3) The sum of 4 skinfolds (Sum 4 skf, mm) was used as an index of total subcutaneous fatness.

- 4) Finally, regional fat distribution was assessed by the trunk-to-extremities ratio of skinfold thickness: $TER = (Sca + Sil) / (Tric + Bic)$ (Malina and others 1999).

In addition, in mothers, body composition was quantified using bioelectrical impedance analysis (BIA). BIA measures resistance and impedance properties of the body. It was introduced and validated several years ago (Foster and Lukaski 1996). We used a Maltron analyzer® (BF 900, Maltron International Limited, Rayleigh, Essex, UK). BIA was performed on the left side of the body by placing 4 electrodes at the sites indicated by the manufacturer (hand-wrist, foot-ankle) (SL & RL). Measurements were taken with the subject lying down, barefoot, and corrected for stature. Values were derived from the prediction equation provided by the apparatus manufacturer for sex, age and geographic origin of the subject. In this report we used percent of fat mass (%FM) as an index of adiposity.

Overweight and obesity were estimated in preschool children according to W-H z-scores (Martorell and others 1998). We used a z-score value >1 to define overweight and >2 to define obesity. The body mass index (BMI, kg/m^2) was computed as $\text{weight} / (\text{height})^2$, and overweight and obesity were also assessed on the basis of International Obesity Task Force (IOTF) criteria (Cole and others 2000). Sex and age-specific cut-off points corresponding to a BMI of 25 kg/m^2 (overweight) and of 30 kg/m^2 (obesity) at 18 years of age were calculated. In mothers, overweight was defined as $BMI > 25 \text{ kg/m}^2$ and obesity as $BMI > 30 \text{ kg/m}^2$.

Stool parasitism

Each child was asked to provide a stool sample in a plastic bag. About 77 % (n=350) of the children returned the bag. A portion of stool was examined in saline immediately after emission. Then the stool was preserved in a formalin solution and carried in a small plastic container to the

SELADIS parasitological laboratory in La Paz. An enrichment procedure with formol-ether-acetate after centrifugation was performed in order to observe the concentration of helminth eggs. Slides were read twice by 2 different observers (SL & SR) and strong agreement was found (>90%). In the present study, the presence or absence of soil-transmitted nematode eggs after concentration (mainly *Ascaris lumbricoides*, *Trichuris trichura* and hookworm) served as an overall index of parasitism in children. Presence of helminths in children suggested a low level of hygiene in the familial environment (codification: 1= no infestation; 2= presence of helminths).

Economic activity, diet and village accessibility

These consisted of information on the main household productive activities (agriculture, fishing, forestry, public office, trade) via questionnaires give to the mothers (n=172). Family economic activity was encoded into 2 categories. 1: agriculture, fishing and wood cutting; 2: wage earning, storekeeping and handicraft. Mothers were also asked how they self-classified into an ethno-linguistic group (Takana, Esse-Ejjas, Aymara, Quechua, and Castellán). This variable was used as a proxy for the cultural background of the household. Only 172 mothers (88%) responded correctly to the questionnaire after careful checking. Twenty-three mothers gave incomplete responses, mainly because they did not have enough time to answer, and their responses were discarded for analysis. One hundred forty-eight families out of 172 (86%) self-considered as Takanas; 23 (13%) as Esse-Ejjas and 1 as colonist Quechua-Aymara.

An investigator trained in nutrition (SL) administered a questionnaire on dietary habits to the mothers. They first responded to a 24 h food intake recall, recapitulating the main dishes consumed in the household. Next they were asked about frequency of foods consumed on a daily and weekly basis. Food items were allocated into 7 groups: fish, cereals, tubers, plantain, fruits, meat and milk. Fish and plantain constituted the main caloric source of the diet. A food diversity index (FDI) was elaborated, recording the presence or absence of consumption of food in each group. This index was derived from a diversity score developed in West Africa and adapted to the present situation (Savy and others 2005). During our visits, there were no evidence of energy shortages, but the overall nutritional value of diet appeared suboptimal. This is a source of concern, since nutrients such as the essential amino acids, zinc, iron, calcium and some vitamins are crucial in preventing stunting (Golden 1991). Hence, it was decided to introduce weighting to

confer more importance to group food with higher nutritional value. This weighting gave a higher rate to foods of animal origin, then to cereals and fruits, and finally a lower value to starchy foods like tubers or plátano. Four categories were computed:

- 1) category 1: fish, meat and milk (weight=4)
- 2) category 2: cereals (weight=3)
- 3) category 3: fruits (weight=2)
- 4) category 4: tubers (mainly yucca or *Manihot esculenta*) and plátano (*Musa paradisiaca*) (weight=1)

Interestingly, the FDI had a normal distribution (mean: 10.7 ± 3.2 ; median : 11; min: 4; max: 18).

The median of FDI, equal to 11, was used as a cut-off point for dietary diversity (code= 1: high diversity (FDI>11); code=2: low diversity (FDI<11)).

Finally, village accessibility, which conveys the idea of better access to health facilities or market foods for communities living near the district town, was coded into 3 classes:

- 1) Easy to access all year round by road or canoe in no more than 2 h
- 2) Accessible by canoe only, within 2 to 5 h downstream
- 3) Accessible by canoe only, within 6 to 12 h downstream

Analysis

Data adjustments

Normality of variables was checked and could be assumed for anthropometric and BIA measurements, but not for subcutaneous thickness. Skinfold distribution was heavily skewed to the right. A log natural transformation was thus applied to normalize it. In mothers, BMI distribution was slightly asymmetric. It was normalized after a log transformation. Log naturals of the sum of 4 skinfolds (log sum 4 skf) and of TER (log TER) were used for analysis. During childhood, body composition measurements vary strongly with age. To control this age effect, stepwise regression between the different indices (BMI, S4skf, TER and AMC) and polynomial expression of age (age, age² and age³) were performed (p to enter <0.05). The residuals from the regressions were then retained to represent the different body composition indices adjusted for age effects.

2 *Statistics*

4 Continuous variables were expressed as means and standard deviation (SD). The 95%
confidence interval was indicated in case of prevalence estimates. Bivariate comparisons were
6 performed with the Student t test or the χ^2 test in case of categorical variables. Multiple
associations between the different indices of body composition were studied with multiple
8 regression procedures. Residuals on age of anthropometric measurements, instead of crude
values, were entered into the regressions to control for an age effect. Multicollinearity of
10 variables was checked. Finally, a logistic regression was run to analyze the influence of
categorical variables (familial variables, health status, parasitism or FDI) on the risk of
12 overweight. Data entry and clearance were performed using Epi Info software (2004). Statistical
analysis was done using SPSS 11.01 (<http://www.spss.com/>) and NCSS software
14 (<http://www.ncss.com/>). The level of significance was set at $p < 0.05$.

Results

Chronic malnutrition and overweight were estimated on the basis of anthropometric indices. According to these indices, 39.4% (IC_{95%}: 34.0~44.0%) of the 0-10 year-old children were stunted in growth (H-age <-2 z-scores) and 9.2% (IC_{95%}: 6.3~12.7%) presented a weight deficit (W-age <-2 scores). No children were wasted. Among mothers, only 2 had a BMI between 17.5 and 18 kg/m², (entry point to chronic energy deficiency). There was no sex difference in the W-H index; hence, values were combined. Figure 2 indicates changes in W-H with age. Values were above the references at all ages. A trend toward a linear increase with age was observed. The W-H index could not be computed in boys taller than 145 cm or in girls taller than 137 cm; BMI values were used instead. Contrary to W-H, there existed significant sex differences in BMI. Curves were thus drawn separately (Figs. 3 a-b). Curves from both sexes suggested a curvilinear relationship, with a minimum reached by 7 years of age. Values were higher than the 50th percentile of the reference for age and sex up until 10 years of age. Thereafter, differences were less noticeable in boys.

Overweight in preschool children was estimated using z-score cut-off values: about 17 % could be considered as overweight (Table 1). IOTF criteria were used in children older than 2 years. As a whole, 12.2% of children were overweight and only 1% were classified as obese. Differences existed between age groups: there were significantly more overweight preschool children (20.9%) than older children: 8.5% in the 5-10-year-old group and 9.6% in the 10-14-year-old group ($\chi^2=7.6$, DF=2, $p<0.02$). In women, BMI cut-offs of 25 and 30 were used as limits. A total of 35.3% of women were overweight but only 5% were overtly obese. There was a non-significant trend toward an increase with age: 25% in women under 25 years of age compared to 42% in women older than 35 ($p=0.08$).

Regional fat distribution, judged from the TER, varied with age. There was more fat trunk accumulation in adolescents (TER > 1; TER= 1.30 ± 0.29) than in preschool (TER= 0.96 ± 0.20) or school-age children (TER= 1.09 ± 0.20) ($p<0.001$). In women, the ratio was high (1.61 ± 0.35). Older women had a higher ratio than younger mothers. Normally nourished children had significantly greater subcutaneous fat mass than stunted children, but no difference was apparent in TER until 10 years of age, at which time normal adolescent males (not females) had higher fat trunk deposits than stunted males ($p<0.04$).

Because of the short stature of this population, it was of importance to examine relationships between BMI and other indices of body composition or regional fat distribution. Hence, multiple regression analysis was carried out between the BMI (dependent variable), AMC (muscularity indicator), log sum 4 skf (fatness indicator) and log TER (fat distribution indicator) in children (Table 2). Analyses were run separately in boys and girls to take into account sex differences. As stated above, residuals on age were used to control for an age effect. In boys, body composition variables explained 43% of the variance in BMI ($p<0.001$); AMC and the log sum 4 skf had significant and positive effects, but log TER had no influence. In girls, the model explained a higher percent of BMI variance (52%, $p<0.0001$). All variables were significant predictors. Interestingly, log TER had a negative effect on BMI: a lower ratio was associated with higher BMI. In women (Table 3), in addition to the previous variables, the total body fat mass calculated after bioelectrical impedance measurements (TBF) was added. The model explained a higher proportion of BMI variance (86%) than in children. All variables apart from TER had a positive and significant effect on the variance of BMI.

In order to analyze cross-relationships between ethnic origins, family activity, health status, dietary diversity, parasitic infestation and village accessibility and the risk of overweight, children were separated into 2 groups, normal and overweight, according to IOTF criteria. We used cut-off points presented in the Methods section. We were unable to find any significant differences between normal and overweight children with respect to these variables. The same analysis was repeated in women. They were classified as normal ($BMI<25 \text{ kg/m}^2$) or overweight ($BMI \geq 25 \text{ kg/m}^2$). In contrast to children, all cross-relationships between familial and environmental variables and the risk of overweight were significant. Village accessibility was the unique exception. To further explore these relationships, logistic regression analysis was run between categories of BMI and the previous categorical variables (Table 4). The model was pertinent ($\chi^2=27.3$, $DF=4$, $p<0.0001$) and correctly classified 67% of the subjects. Both FDI and health status score had a significant effect: a low FDI and/or mediocre health status were associated with lower BMI. However, ethnic group and economic activity variables no longer had a significant independent effect.

Finally, the issue of the relationship between body composition indices of children and BMI of their mothers was examined using multiple regression procedures. The model was adjusted for

the age effect using residuals on age of the different indices of body composition. Regressions
2 were run separately in girls and boys. Results are shown in table 5. In boys, predictions of total
fatness and arm muscle circumference based on maternal BMI were significant. However, in
4 girls, maternal BMI did not predict any of the body composition indices.

6

Discussion

Results of the study on this particular population confirm the existence of an overweight risk in women and perhaps in children, while a substantial percentage of children (39.4%) still suffer from chronic malnutrition (stunted growth). However, the overt obesity prevalence remains low. Indeed, compared with the most recent DHS survey, these values are apparently lower. As a whole, 35.3% (IC_{95%}: 28.7~42.5%) of women were overweight or obese compared to 45.5% found at the national level and 50.7% in the Amazon departments of Beni and Pando (MEASURE 2004). No recent estimates exist for children, but the average W-H was 0.4 z-score in 2003 and 0.2 z-score in the present survey. Martorell and co-workers (Martorell and others 1998), from DHS 1994, found an overall prevalence of overweight in 15.2% of the 1-3-year-old children. This figure very closely resembles that of the present study, 17 %.

Indicators of overweight or fatness based on weight and height have been consistently criticized because they might be equally influenced by fat mass and fat-free mass. Limitations to the BMI in children have been stressed (Malina and Katzmarzyk 1999). There exist important differences between sexes, ethnic groups and age groups, resulting in poor BMI sensitivity when classifying fatness (Baumgartner and others 1995). Despite these limitations, BMI is now considered an international reference for overweight in children (Cole and others 2000; Dietz and Bellizzi 1999). In that study, the BMI was only moderately related to other body composition indices in children. This suggests that the BMI might be better viewed as an index of heaviness (Campbell and others 2001). BMI was inversely correlated with TER in girls, indicating more fat accumulation in the limbs. This finding points to sexual dimorphism in overweight girls, suggestive of gynoid deposition (He and others 2004). In contrast, in women, BMI was strongly linked to subcutaneous fatness. In general, there was an increase in BMI and W-H with age.

Familial clustering of overweight is an important issue within the framework of the “nutritional transition process” in which changes in food patterns and lifestyle occur very quickly (Popkin 2001). One might then ask whether recently acquired fatness traits and their concomitant comorbidities will be transmitted to the next generation. Indeed, familial resemblance in total fatness and regional fat distribution have been reported (Li and others 1996). Moderate familial

similarities between mothers and children were found in the Oaxaca Valley of Mexico (Little and Malina 2005). In the present study, the prediction of children's phenotypes from maternal BMI was moderate in boys and practically non-existent in girls. Body composition characteristics of mothers and children seem disconnected. It is thus tempting to infer that the risk of overweight in mothers and children is of a different nature and originates from different factors. It is conceivable that the high prevalence of overweight in mothers is the consequence of recent events in terms of changes in food patterns or physical activity. These changes seem to only slightly affect children, perhaps because of persistence of a heavy and widespread burden of infection and poor hygiene status, evidenced by severe intestinal parasitism. This is in line with the coexistence of overweight in mothers and stunting in children which was observed in populations in transition (Garrett and others 2005).

The existence of a link between stunting and fatness was postulated in previously malnourished children from Brazil (Sawaya and Roberts 2003). Likewise, several observations from Latin America have confirmed the association between early malnutrition, represented by short stature, and later adiposity (del Rio-Navarro and others 2004; Doak and others 2002; Garrett and others 2005; Schroeder and Martorell 1999; Sichieri and others 2003). In the present study, there was no association between actual stunting in growth and overweight or regional fat distribution. While the transversal nature of the data collected does not permit such a demonstration, it also does not negate it. Likewise, a phenotype associating short stature with normal weight or slight overweight in children has already been described in other Amazonian communities (Benefice and Barral 1991; Capelli and Koifman 2001; Mattos and others 1999; Santos and Coimbra Junior 1991), and in the Bolivian Amazon (Foster and others 2005).

A tendency to accumulate fat in populations experiencing agricultural or economic development is likely in the Amazon (Gugelmin and Santos 2001). The present findings shed light on more specific factors. Body composition indices observed in children and adolescents were not influenced by factors such as health status or parasitism, nor by dietary intake. It is possible that, in the particular context of this survey, the different components of growth (linear growth and body composition) do not respond in the same way to environmental or familial factors. However, it must be underlined that body composition of children is mainly dependent on biological and physiological factors like the tempo of growth and maturation before and during puberty (Malina and others 1999). In contrast, in mothers, the risk of overweight was

significantly related to food diversity and health status. In its elaboration, the FDI assigns a greater value to food of animal origin. Hence, women who ate more fish, meat, bread and rice had a higher FDI and were more prone to overweight than mothers relying mostly on starchy foods like cassava roots and plátano, maintaining a normal BMI. Thus, the FDI could reflect two different ways of life in the area: the subsistence level of small communities of farmers and fishermen dependent on their own production and consuming a monotonous diet, and a population with agricultural and forestry activities integrated into a larger economic market.

The challenging aspect of the present survey lies in the fact that it was performed in an isolated rural area where overweight is likely to be infrequent because of the level of physical activity needed to cope with everyday duties (farming, fishing, carrying water, collecting plants, hunting, cutting wood, etc.) and the persistence of food instability during drought periods (Dufour 1984). This would signify that the nutritional transition theory designed by Popkin (Popkin 2002) is not confined to urban or peri-urban areas of developing countries, but is evenly scattered. In many parts of the Amazon, including the Beni, important economic changes are in progress due to the building of roads and the development of forestry and tourism. Formerly isolated populations are now incorporated into a large market exchange and in the monetization process that may have both positive and undesirable effects upon health and nutrition (Leatherman and Goodman 2005 35). Health implications of this phenomenon are not clearly elucidated in the area studied. Given its importance in other parts of Bolivia (Barcelo and others 2001; Perez-Cueto and Kolsteren 2004), it is likely to be harmful and warrants high priority from health institutions.

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Figure 1: Map showing the location of the different communities (after GPS reading)

Figure 2: W-H for age in children

Figure 3 (a-b) BMI for age in boys and girls

Table 1: Prevalence of overweight in Beni riverside children and women

Age group	Criterion	Overweight (%)	Obesity (%)
0-5-years-old (n=171)	W-H z-scores	17.0 ¹ (10.0~21.0) ²	3.0 (0.3~5.0)
2-5-years-old (n=105)	IOTF ³	20.9 (13.6~29.9)	2.8 0.6~8.0
5.1-10.0-years-old (n=176)	IOTF	8.5 (4.8~13.6)	0.5 0.1~3.1
10.1-15.0-years-old (n=104)	IOTF	9.6 (4.7~16.9)	0.0 (0.0~3.0)
All ages (n=385)	IOTF	12.2 (9.1~15.9)	1.0 (0.3~2.6)
Women (n=195)	BMI cut-offs	35.3 (28.7~42.5)	5.0 (2.5~5.0)

¹: Prevalence (%)

²: Lower and upper 95% confidence interval

³: International Obesity Task Force

Table 2 Relationships between BMI and body composition variables

1 a: Boys (n=238)

$R^2 = 0.43$; $F = 58.9$, $p < 0.0001$

Independent variable ¹	Regression coefficient	Standard error	t value	p
Intercept	0.05	0.07	0.78	0.44
AMC	0.50	0.06	8.38	0.00
Sum 4 skf	2.10	0.26	7.97	0.00
TER	0.33	0.37	0.90	0.37

¹ Residuals of arm muscle circumference, log sum 4 skinfolds and log trunk-to-extremities ratio on age

1 b: Girls (n=212)

$R^2 = 0.52$; $F = 77.3$, $p < 0.0001$

Independent variable ¹	Regression coefficient	Standard error	t value	p
Intercept	0.06	0.07	0.92	0.36
AMC	0.43	0.05	8.02	0.00
Sum 4 skf	3.17	0.30	10.66	0.00
TER	-0.97	0.37	-2.60	0.01

¹ Residuals of arm muscle circumference, log sum 4 skinfolds and log trunk-to-extremities ratio on age

Table 3: Relationships between BMI (log natural) and body composition variables¹ in women (n=191)

$R^2 = 0.85$; $F = 259$, $p < 0.0001$

Independent variable ¹	Regression coefficient	Standard error	t value	p
Intercept	1.84	0.05	31.9	0.00
AMC	0.02	0.002	10.7	0.00
Sum 4 skf	0.13	0.016	8.33	0.00
TER	-0.006	0.018	-0.33	0.74
TBF ²	0.0093	0.001	9.01	0.00

¹ Log natural of sum 4 skf and trunk-to-extremities ratio

² Total body fat (after BIA)

Table 4 Summary of a logistic regression analysis between overweight risk and familial or individual characteristics in women (n=172)

	Regression coefficient (B)	Standard error	Wald test	ddl	p	Odds ratio (Exp (B))
Ethnic group	-7.69	13.6	0.31	1	0.57	0.00
Economic activity	0.64	0.6	1.15	1	0.28	1.9
Health status	-0.65	0.34	3.68	1	0.05	0.51
FDI	-0.74	0.34	4.60	1	0.03	0.47
Intercept	8.71	13.7	0.40	1	0.52	

R^2 Cox & Snell= 0.14; χ^2 model= 27.3 (DF=4, $p<0.0001$)

Table 5: Prediction of body composition indices¹ of children based on maternal BMI

Table 5 a: Boys (n=238). $R^2=0.12$ (F=8.6, $p<0.0001$)

Independent variables	Regression coefficient	Standard error	t	p
Intercept	24.4	0.19	126.7	0.000
AMC	0.48	0.20	2.3	0.01
BMI	-0.33	0.19	-1.7	0.08
Sum 4 skf	4.4	0.87	5.0	0.000
TER	0.48	1.08	0.4	n.s

Table 5 b: Girls (n=212). $R^2=0.03$ (F=1.82, $p=0.12$)

Independent variables	Regression coefficient	Standard error	t	p
Intercept	24.4	0.21	112.8	0.000
AMC	-0.18	0.18	-0.9	n.s
BMI	0.05	0.21	0.27	n.s
Sum 4 skf	2.04	1.12	1.8	0.07
TER	0.77	1.15	0.6	n.s

¹ Residuals of AMC, BMI, Sum 4 skf and TER on age